

## Primacy and Economic Development: Bell Shaped or Parallel Growth of Cities?

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The share of the largest city in the total urban population - the so called primacy ratio - differs substantially across countries. One hypothesis in the literature is that economic development first increases and subsequently decreases primacy. The alternative hypothesis is that cities grow in a parallel way, such that primacy ratios depend on geography or history rather than economics. I test these hypotheses for a cross section of countries. Evidence is found for a bell shaped relationship between primacy and economic development conditional on the potential of countries to develop a balanced urban system. Still historic variables influence current primacy ratios.

### I. Introduction

The share of the largest city in the total urban population - the so called primacy ratio - differs substantially across countries. The ratio is, for instance, 6 percent in India, 38 percent in Argentina, 33 percent in Korea, and again only 8 percent in the Netherlands (see Table A1). Development and urban economists have long tried to link primacy and economic development. One hypothesis is that economic development first increases and then decreases primacy or other measures of spatial concentration within a country - thus exhibiting a bell shaped relationship (Alonso (1980), Williamson (1965)).

An alternative hypothesis is that cities grow in a parallel way, which means that spatial concentration is unaffected by urbanization and economic development (Black and Henderson (1997) and Eaton and Eckstein (1997)). Then, the distribution of the urban population may simply reflect geography or historic shocks, which created a functional system of cities and cemented a size distribution, such that further urbanization only increased the population of cities in a parallel way.<sup>1</sup> A persistent influence of ancient population distributions would point to the importance of path dependencies, but also to ongoing natural advantages of certain locations in a country. This alternative hypothesis is far less appealing, because it does not promise any relief from the present problems that several developing countries have with regional disparities, congestion and the excessive growth of their primate cities. Therefore, considerable

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1. For instance, the fact that Vienna by so much dominates the urban system of Austria is more likely to be traced back to its century long role as a capital of a huge empire than to current GDP per capita levels. Similarly, France is more concentrated than Germany, because it has no federal political history but has been centrally governed from Paris for the longest part of the recent history.

attempts have been made to reveal the bell shaped relationship of primacy and economic development in time series and cross country analyses. While time series analyses often find such a pattern, cross country evidence is mixed.

This paper makes a new attempt to show the existence of the bell shaped relationship in a cross country setting, using different country samples and a number of conditional variables that influence the likelihood that a country develops a dispersed or a concentrated urban system. The conditional variables are the population density, the size of the country, the transportation infrastructure and political variables like openness, political and economic freedom and a colonial heritage.

I first review the existing empirical literature in Section II. Section III reviews the theoretical arguments for the existence of the bell shaped or inverted U-curve and other determinants of urban concentration. Section IV discusses the empirical procedure and presents the results.

## **II. Previous Empirical Evidence for the Bell Shaped Curve**

*Time series studies* find ample evidence for the bell shaped curve in single countries. El-Shakhs (1972) shows the existence of the curve for the UK, and Alperovich (1992) shows its existence for Israel. Parr (1985) estimates Pareto coefficients for 12 countries and several years between 1850 and 1981.<sup>2</sup> He finds strong evidence for the curve in the high income countries Austria, France, Sweden and the US. Concentration peaks around 1910 in Austria, 1930 in the US and Sweden, and 1954 in France. The evidence for the curve is weaker for Brazil, Japan, Spain and the USSR, while in the low income countries Egypt, India, Nigeria and Turkey no clear reversal of the trend towards increasing concentration can be observed yet.<sup>3</sup>

*Cross-country studies* find mixed evidence for the bell shaped curve. Williamson (1965), Kamerschen (1969), El-Shakhs (1972), Wheaton and Shishido (1981), and DeCola (1984) find evidence for the curve between urban concentration and economic development. Kamerschen (1969) uses data for 80 countries and the share of the largest city in percent of the four largest cities as a measure of concentration. He finds a negative relationship between urban

2. The estimation of Pareto coefficients goes back to Singer (1936). For a "Pareto coefficient" of 1, one gets the so called rank size rule, which states that the rank times the population of a city equals a constant, which is the size of the largest city. This "rule" has been refuted by Hsing (1990) and Alperovich and Deutsch (1995).
3. Eaton and Eckstein (1997) calculate Lorenz curves for city size distributions for France since 1876 and Japan since 1925. In contrast to Parr's evidence, these authors argue that relative city sizes in France and Japan did not vary that much to exclude the possibility of parallel growth in the course of economic development. See Carroll (1982) or Kasarda and Crenshaw (1991) for a survey of further early time series and cross section studies on the distribution of city sizes. For an overview of determinants of urban concentration see Mutlu (1989) and Sheppard (1982). Mutlu (1989) presents a number of variables used in previous studies, but finds only weak evidence for an impact of economic variables on concentration measures. GNP per capita reduces concentration in his sample of up to 90 countries. The size of the absolute urban population and absolute total population, as well as the size of the arable land reduce concentration while greater income inequality increases concentration measures. Mera (1973) finds a positive relationship between concentration and development in a sample of 46 developing countries. However, he uses a different dependent variable, i.e., the share of the largest city in the total population instead of in the urban population. This share naturally increases with urbanization and, therefore, is a less measure of spatial concentration.

concentration on the one hand and GDP per capita (GDPC) and industrialization on the other hand for developing countries, and a positive relationship for developed countries. Wheaton and Shishido (1981) test the bell shaped relationship against a linear relationship. Using a sample of 38 countries, that have at least three metropolises, they find that the bell shaped curve much better explains the data.

Rosen and Resnick (1980), Mutlu (1989), Lemelin and Polèse (1995), and Moomaw and Shatter (1996) find a negative relationship between spatial concentration and economic development. Rosen and Resnick (1980) use a sample of 44 developing and developed countries. They test for the significance of several variables on calculated Pareto coefficients. Higher GNP per capita and total population lead to a more evenly distributed population. The density of the rail-network, the export to GNP ratio, the percentage of the non-agricultural labor force as well as a dummy for former colonies are not found to be significant in the overall sample. In a sample of 61 countries, Lemelin and Polèse (1995) find that primacy is indirectly linked to the level of economic development. Primacy falls monotonically with the degree of development. They mention, albeit do not report, that their results are robust to alternative measures of primacy and development, and to alternative country subsets. Moomaw and Shatter (1996) use panel data for 90 countries and find that primacy ratios increase if the largest city is additionally the capital of a country. They also find that GDPC, literacy, population and export orientation reduce primacy. Ades and Glaeser (1995) show the impact of political factors on primacy ratios for 85 countries. Openness and high government expenditures for transportation decrease primacy ratios, while dictatorships and political instability increase them. Primacy is also higher for countries, in which the largest city is additionally the capital.

Richardson and Schwartz (1988) find no support for the economic development and primacy link at all. For the 116 countries under study they show that demographic factors are more important and render economic variables insignificant. They show that 40 percent of the variation of primacy can be explained by national population, urban population and a Latin America dummy. Their results are criticized by Lemelin and Polèse (1995) partly for not testing for multicollinearity. Urbanization is strongly correlated with economic development and GDPC levels. Hence, insignificant coefficients for GDPC levels in the presence of urbanization or industrialization variables should not come as a surprise. Another problem might be the large sample size, which probably includes a large number of very small countries, where no economic rationale predicts systematic domestic economic forces to be able to unfold and to play a dominant role for the city size distribution.

Arbitrary sample selection is often the reason why the previous empirical results are very sensitive to the group of countries included in the analysis. Therefore, the economic rationale for the selection of countries and explanatory variables is thoroughly discussed in Section III and Section IV.2.

### **III. Explaining the Bell Shaped Curve and Further Determinants of Primacy**

The review of the empirical literature of cross-country studies shows that the evidence is mixed as to whether a bell shaped curve exists or not. Thus, the current status and the appropriate attitude towards that curve is somewhat similar to what Williamson (1997) writes

## JOURNAL OF ECONOMIC DEVELOPMENT

about the Kuznets Curve: “now you see it, now you don’t. The important inference of that fact, however, is *not* to reject the Kuznets Curve, but to ask why we sometimes see it and sometimes not.” Most important for the understanding of changes in the concentration and dispersion of economic activities over time is to understand the interplay between centrifugal and centripetal forces that drive this process. In general, positive economies of scale foster agglomerations and industrial clusters, and negative spillovers and higher factor costs foster population dispersion. From that, some authors have derived theoretical explanations for the bell shaped curve.

Wheaton and Shishido (1981) argue that economic development increases capital intensity in industrial production. As capital intensity increases fixed costs compared to variable costs, scale intensity increases. This favors larger cities. Therefore, efficient city size increases with economic development until some sort of capital saturation sets in as scale economies are not exploitable without bound. This again levels the population concentration in later stages of economic development.

Parr and Jones (1983) suggest a five stages approach of economic development and primacy. The pre-urban stage is characterized by a low quality transportation system, which limits the extent of regional markets, intraregional trade and the exploitation of scale economies. In a later stage, improvements in the transportation system allow for more intraregional trade, which pushes some sites above the critical mass of production. This leads to rapid growth of certain specialized cities. In a third stage stronger interindustry linkages further allow exploitation of scale economies in larger cities. In the fourth stage improvements of interurban transportation systems and a high income elasticity of land demand lead to suburbanization and levels economic conditions between cities. This development is extended in the last stage, where several regional markets achieve a sufficient size for the production of a large number of goods. Together with negative externalities from concentration in the core, this leads to a leveling of the population distribution within a country. Thus, falling transport costs have ambiguous effects. They strengthen centripetal forces at low income levels and centrifugal forces at high levels.

Krugman (1991) has formalized the impact of transport costs on the pattern of production. His economic geography model shows the tension between centrifugal and centripetal forces. With high transport costs firms become tied to markets, while for intermediate transport costs, they matter less than the production costs differences.<sup>4</sup> For zero transport costs, space does not matter. Hence, his model predicts industry dispersion for high transport costs, concentration for intermediate transport costs and again dispersion for zero transport costs. This leads to a bell shaped curve between concentration and the level of transport costs.

Krugman’s model has been augmented for the effects of economic development by Junius (1996). Core regions benefit from the sectoral change to industrial production that usually accompanies economic development. As a consequence, a higher share of the work force starts working in the footloose industrial sector under increasing returns to scale. The demand from these workers reinforces centripetal forces that make it even more profitable for firms to cluster. Increasing concentration, however, also leads to negative externalities due to crowding and

4. Due to home market effects, production costs are always lower in the larger region.

## JUNIUS: PRIMACY AND ECONOMIC DEVELOPMENT

pollution. These congestion effects decrease the advantages of being in the core. This slowly levels economic conditions between core and peripheral regions. Thus, peripheral regions gain after some level of concentration is reached in the core. The resulting bell shaped curve between spatial concentration and economic development is intensified by a fall of transport costs.

Economic factors can only play an important role if they are able to unfold themselves without being dominated by other factors. Neighboring countries and geographic or demographic factors may especially determine the population distribution of small countries like Luxembourg or Singapore, where two independent metropolises additionally would not be distinguishable statistically. Also internal transport costs do not matter much in small countries such that any location would have an advantage for just in time production. Small countries are unlikely to form a functional system of cities and pattern of industry distribution independent from their neighbors and geographic factors.

In countries with a low absolute (urban) population the distribution of the population is likely to be determined by endowment factors. Neither positive external economies of scale nor negative external effects of population concentration are strong enough to determine the population distribution between its different parts. This means that congestion effects are unlikely to push people out of Iceland's Reykjavik with 145,000 inhabitants. Iceland's population distribution is more likely to reflect resource endowments and climatic conditions than scale economies. That is, in order to test for the influence of economic determinants on primacy, only these countries can be used that are not too small in terms of the area or the number of inhabitants.

Concluding from the above models, one would expect a positive coefficient of GDP per capita and a negative coefficient of the squared GDP per capita value on a measure of spatial concentration, provided that countries are equal in all other respects besides GDP per capita. However, countries differ in several respects that influence the likelihood of developing a balanced urban system other than the pure economic forces described by the models. To correct for the different potential to develop several urban agglomerations, the degree of concentration should be made conditional on the following variables.

The extent of the transportation system as indicated, for instance, by road kilometers per land area (*ROADLAND*), has an ambiguous effect on the degree of concentration. As economic geography models show, on the one hand lower transport costs increase the degree of competition that peripheral firms face from firms in the core region. This reduces the attractiveness of peripheral sites and increases the advantages of the primate city as a location of production and consumption. On the other hand, lower transport costs mean that distance is getting less important, such that the need to cluster in order to realize scale economies is less severe. Then industries are able to escape higher land prices and congestion effects by shifting production away from the core and still benefit from agglomeration and scale economies. The coefficient of the variable measuring the extent of the transportation system may, therefore, be interpreted as indicating which effect prevails.

Densely populated countries (*DENSE*) are likely to be less concentrated for two reasons. First, along with population concentration come negative externalities like congestion effects and pollution. They constitute an upper bound to city size after which no or only few scale economies can be realized and, thus, provide strong incentives to disperse. Second, high densities

## JOURNAL OF ECONOMIC DEVELOPMENT

provide the possibilities to disperse since a large number of workers will be available in several parts of the country such that scale economies can be exploited not in the primate city only. They constitute the critical mass of workers and local demand for the production of certain goods that is needed for the formation of a city. Thus, through the existence of a lower and an upper bound to city size, densely populated countries are likely to exhibit several “optimal sized” cities. This reduces urban concentration.

A large land area (*LAND*) increases the probability of forming several metropolises, and, thus, leads to lower levels of concentration in a country.<sup>5</sup> First, it increases the number of possible sites for potential cities that emerge because of the availability of natural resources or natural advantages like ports or transportation nodes. Second, it implies large distances between different parts of a country. Such differences may have favored the development of different urban systems. This is, because historically, the extent of the market for perishable goods has been much more limited than it is today. Additionally, certain services and administrative functions had to be located close to the generally dispersed rural population, such that large countries developed a larger number of medium size cities.

In addition to these variables indicating the possibilities of a country to disperse, the degree of concentration can be influenced by politics.<sup>6</sup> An open trade regime (*OPEN*) favors export industry and population dispersion, because being close to the center of the home market becomes less important. Instead, being close to the market of the neighboring countries, to transportation nodes and harbors becomes decisive. This disperses industries and population. Openness also means better opportunities for non-industrial activities like agriculture. Agriculture is often discriminated in less open regimes, which prevents the development of dispersed food processing and other rural industries.

Undemocratic institutions, the deprivation of civil or political rights, of property rights and domestic unrest (*POLITICS*) are likely to favor urban concentration. In countries with these properties, there often is a strong central and primate bias of government spending. This reflects that in such political systems, spatial politics are often used to assure maximum political control over a country, population and administration. Government expenditures subsidize services to the favor of the urban class which is the backbone of political support. Together with a tendency to nepotism in government spending, this increases the size of the primate city.

Finally, a colonial history is likely to favor strong primate cities. For administrative reasons, the capital of a colony had strong links with the colonial power and, consequently, less strong links with the rest of the national urban system. Production and trade was often more oriented towards the demand of the colonial power than towards the demand of the domestic population. Thus, it often took place in export exclaves, which delinked the capital from the rest of the country and prevented the development of a dispersed domestic urban system. That also prevented the establishment of a traffic network with neighboring countries.

5. Note the economic difference between the variable *LAND* and *DENSE*, although the construction of *DENSE* requires use of *LAND*. The coefficients of correlation between the variables varies between -0.23 and -0.45, depending on the sample used in the following empirical analysis.

6. Krugman and Livas Elizondo (1996) have extended the model for another region called rest of the world. They showed that open trade policies towards the rest of the world favor domestic industry dispersion.

In addition, innovations and other positive growth shocks diffused more slowly to secondary centers as they did in countries without a colonial history. The result is a persistent dominance of the capital city with few and much smaller rival cities. The dominance is expected to be the stronger, the longer countries have been under colonial rule. Therefore, depending on the date of independence, two colonial dummies are used, to approximate their different historic or initial conditions.

#### IV. Empirical Results

##### 1. Estimation Equation

A bell shaped relationship between primacy and GDP per capita (GDPC) exists if primacy depends positively on GDPC and negatively on the squared GDPC value, holding constant the conditional variables. Hence, I estimate variants of the general equation

$$PR = \delta_0 + \delta_1 RGDP C + \delta_2 RGDP C^2 + \delta_3 X + \mu, \quad (1)$$

where  $PR$  is the primacy ratio,  $RGDP C$  is PPP adjusted GDPC,  $X$  is a vector of further explanatory variables,  $\delta_0$ ,  $\delta_1$ ,  $\delta_2$  and  $\delta_3$  are the coefficients to be estimated, and  $\mu$  is the error term. Since primacy measures are expressed in percentage points and, hence, can only vary between 0 and 1, they are limited dependent variables. Thus, to avoid estimation and interpretation problems, I assume that  $PR$  is distributed according to the logistic function. That is:

$$PR = \frac{e^z}{1 + e^z}, \quad (2)$$

where  $z = \delta_1 RGDP C + \delta_2 RGDP C^2 + \delta_3 X$ . By the use of (2), primacy can be transformed into a variable that is not limited to the 0 to 1 range anymore:

$$\frac{PR}{1 - PR} = e^z. \quad (3)$$

Taking logarithms on both sides, one arrives at an estimable equation:

$$L = \ln(PR/(1 - PR)) = z + \mu \quad \text{or} \quad (4)$$

$$L = \delta_1 RGDP C + \delta_2 RGDP C^2 + \delta_3 X + \mu. \quad (5)$$

## 2. Data

The sample of countries is selected as follows. I start with all 209 countries listed in the World Bank CD-ROM (1996) and use their 1990 values unless stated otherwise. As explained in Section III, I first exclude all small countries. That is, I exclude countries with a land area below 30,000 square kilometers, which is slightly below the size of Belgium and slightly above the size of Haiti. I also exclude countries with an urban population below 1,000,000 or a total population below 3,000,000. The reason is that small countries' population patterns are very likely influenced by the spatial structures of their neighbors. They are less likely to have developed an own independent spatial system influenced by systematic economic phenomena only, such that a test of an economic theory would not yield interpretable results.

From the remaining countries, 16 countries are excluded because no PPP based GDPC levels are available for either 1989, 1990 or 1991. The Dominican Republic, Myanmar, Saudi-Arabia, Nicaragua and Sudan are excluded because data is missing for other variables. Observations for Togo, Somalia, Sierra Leone, Senegal, Mozambique, Madagascar, Honduras, Guinea, Chad and Angola are excluded because the last census from which urban population data for 1990 is estimated or interpolated is older than 15 years. Further excluded are the observations for Costa Rica and Zimbabwe as they show large differences according to different sources, and Malawi for which recent estimates are only available for the capital but not for the primate city of the country. The remaining sample of 70 countries is listed in the appendix (Table A1).<sup>7</sup>

Data for the size of urban agglomerations is available from the World Urbanization Prospects of the UN (1995a) and relies on data from national sources. I complete the urban data by data from the UN Demographic Yearbooks (1992, 1995b), because the World Urbanization Prospects report urban data only for capital cities and cities with a population of more than 750,000 inhabitants. The advantage of the World Urbanization Prospect Data is that it provides estimates for all countries in a single common year. However, this comes with the drawback that only estimated or interpolated data can be used, which is likely to inhibit measurement errors. Another problem arising from these data is that estimates for urban agglomerations may differ between the two sources. The statistical concept of measuring city size is different in different countries. Some of them report city proper data, others metropolitan area or urban agglomeration data. I use urban agglomeration data where possible, because it better measures the true concentration of a country. On the basis of Equation (4), I construct, four dependent variables (*PRIMA1*, *PRIMA2*, *PRIMA3*, *PRIMA4*) indicating the share of the one, two, three or four largest cities in the total urban population. *PRIMA2-PRIMA4* are frequently used in addition to *PRIMA1*, in order to consider the size distribution of cities below the largest city. The different levels of *PRIMA1* are shown in the Table A1.

The following variables are taken or calculated from the World Bank (1996): Population density per square kilometers (*DENSE*), land in square kilometers (*LAND*), density of the transportation system measured as the ratio of road length to land area (*POPURB*), total population

7. The former CSFR takes the place of the Czech Republic. Data for Czechoslovakia is taken from the World Bank (1996).



## JUNIUS: PRIMACY AND ECONOMIC DEVELOPMENT

(*POPTOT*), and urban population (*POPURB*). For Russia, the surface instead of the land area has to be taken. This includes also the surface of interior seas or lakes. Data on road length for Bulgaria, Poland, Rumania is from 1989, for Hungary from 1988 and for Iran from 1985.

Data for PPP adjusted real GDPC (*RGDPC*) is taken from the Penn World Tables (Mark 5.6). See Summers and Heston (1991) for a description. For Niger, Romania and Russia *RGDPC* data is from 1989. Data for openness (*OPEN*) and the political regime (*POLITICS*) is taken from Sachs and Warner (1995). According to their definitions (pp.10-11), *OPEN* is a dummy variable that takes the value one if “a very high proportion of imports (is) covered by quota restrictions”, “for Sub-Saharan Africa, (if) a high proportion of exports (is) covered by state export monopolies and state-set prices”, the country has “a socialist economic structure”, or if “a black market premium over the official exchange rate of 20 percent or more, on average, (prevailed) either for the decade of the 1970s or the decade of the 1980s”. Politics is a dummy variable that takes the value one if the country has “a socialist economic structure”, “extreme domestic unrest, caused by revolutions, coups, chronic civil unrest, or a prolonged war with a foreign country that is fought on domestic territory”, or “extreme deprivation of civil or political rights” (Sachs and Warner (1995, p.9)).

Data on the colonial history and the date of independence is taken from Fischer Weltalmanach (1997). I use two dummy variables to distinguish different lengths of colonial rule over a country (*COLONY1815* and *COLONY1950*). *COLONY1815* takes the value 1 if the country was a colony in the year 1815. Colonial powers, independent countries and countries with a long urban history, like China, Germany, Iran or Thailand, take the value “0” also if they did not exist as a political entity at that time. *COLONY1950* indicates countries with a very short history of independence. It takes the value “1”, if a country was still a colony in 1950, and “0” otherwise. In case the date of the proclaimed and final or recognized independence differ, I use the year of the proclaimed independence.<sup>8</sup>

Data on historic variables is available from Banks (1971) for 45 countries. I use 1919 data for railroad mileage per square mile (*RAIL1919*), telegraph mileage per square mile (*TEL1919*) and population density (*DENSE1919*). Density data for China is from 1911, for Ireland from 1922, and from Korea for 1904. GDPC data for 1913 (*GDPC1913*) is available from Maddison (1995) for 44 countries. For 26 countries, the share of industry in total industrial and agricultural employment in 1870 is taken from Mitchell (1992, 1993, 1995). The agricultural sector encompasses agriculture, forestry and fishing. The industrial sector encompasses extractive industry, manufacturing industry, construction and services.

Data on historic urban concentration is taken from two sources. For 41 countries data for the size of the largest in percent of the three largest cities is available from Jefferson (1939). Concentration is measured around the year 1935 according to availability. For European countries data for the size of cities for the years 1800 and 1850 is available from Bairoch *et al.* (1988). From this data, I calculate primacy ratios for 1800 and 1850 (*HISTO1850* and *HISTO1800*). It should be noted that the data used for the construction of today’s primacy ratios is based on the population of urban agglomerations, which are likely to encompass

8. I use the years 1815 and 1950, because they are two landmarks in world history. 1815 was the end of the Napoleonic Wars and the Congress of Vienna. 1950 roughly marks the end of World War II and its postwar turmoil.

cities and districts that are not included in the population of Bairoch *et al.* (1988) city proper data for the 19th century.<sup>9</sup>

In order to avoid getting very small regression coefficients, I standardize *LAND* by dividing it by the size of the smallest country, which was Belgium with 30,260 square kilometers. The resulting variable *NLAND* varies between 1 and 564. I standardize *RGDPC* by the value of the poorest country, namely Zaire. The resulting new variables *NRGDPC* and *NRGDPC2* now range from 1 to 46.3 and from 1 to 2,147. I standardize *TEL1919* by the smallest value of the sample, such that the new variable *NTEL1919* varies between 1 and 339.

### 3. Results

As indicated in the last section, existing data for the size of the largest cities is not collected in a standardized way, but with different statistical concepts and from different sources. Additionally, not all explanatory variables are available for the full sample of 70 countries. Hence, I use different country samples to estimate the likely determinants of urban concentration. The results are reported in the appendix (Table A2). The reported standard errors are corrected by White's heteroscedasticity consistent variances and covariances.<sup>10</sup>

I start with a small but most reliable sample of countries. This sample includes 23 countries for which the World Urbanization Prospects reports at least 4 cities with a population of more than 750,000 inhabitants each.<sup>11</sup> For these countries all four primacy ratios can be calculated from the same data source without relying on additional data from the UN Demographic Yearbooks. This sample is not only most reliable in terms of the data. It is also the sample where economic forces have the highest potential to unfold and not to be suppressed by idiosyncratic or genuine influences. This is because this sample only consists of large countries with a large urban population and a system of several cities.

In regression 1, I start with the full set of current independent variables except for *COLONY1950* because the sample does not contain countries that were a colony in 1950. Except for *OPEN*, all variables have the expected sign, but the equation only explains 20 percent of the observed cross-country variation of urban concentration as measured by *PRIMA1*. The coefficients of *NRGDPC* and *NRGDPC2* are statistically significant at the 5 percent level. The coefficient of *NLAND* is significant at the 1 percent level. The coefficients of *OPEN* and *POLITICS* have large standard deviations. t-tests and a redundant variable F-test do not support the hypothesis that trade policy and "bad" politics influence the degree of urban concentration. The standard deviations of *ROADLAND* and *DENSE* are large. Together with a high correlation coefficient between the variables, this points to the plausible multicollinearity between population density and density of the road system.

9. Therefore, I also use the percentage of the largest city in 1850 and 1800 instead of the percentage of today's largest city in total 1800's or 1850's urban population.

10. I also report the White-test without cross terms. It shows the probability of no heteroscedasticity in the data, which is, for instance, 12.6 percent in the first regression. In the presence of heteroscedasticity the coefficients are still unbiased, but not efficient, such that t-test and F-test are not interpretable.

11. These countries are marked with a # in Table A1 in the appendix.

Multicollinearity means that it is impossible to isolate the individual impact of the variables statistically. OLS estimates remain unbiased and consistent, but have larger standard errors. Thus, confidence intervals get larger, so that the sample data does not provide the information required. Therefore, in the following regressions, I exclude one of the variables if collinearity between them is high in that sample and report preferred estimates only. Omitting a variable that should be included on theoretical grounds leads to a specification bias, because the remaining variable then measures the combined effect of the correlated variables. Since I am mainly interested in the coefficients of *NRGDPC* and *NRGDPC2*, I accept this possible bias and focus on the identification of the bell shaped curve.

Regression 2 reports my preferred estimates for the sample of 23 countries, excluding *DENSE*, *POLITICS* and *OPEN*. An F-test shows a probability of 84 percent that the omitted variables can indeed be excluded from the regression. I also use the regression specification error test (RESET) to check whether the structural specification is subject to the problem of omitted variables. It tests the null hypothesis that the expected value of the disturbance term conditional on the regressors equals zero.<sup>12</sup> The probability that this is the case in this regression is 67.7 percent. Additionally, I test the normality assumption of the classical normal linear regression model. I report the Jarque-Bera statistic and the corresponding probability that the residuals are normally distributed, which is 51 percent in this regression. As before, I find that the coefficients of *NRGDPC* and *NRGDPC2* have the expected sign. They are statistically significant at the 1 percent level. The adjusted  $R^2$  rises to 0.30.

In regressions 3-7, I test whether the results are sensitive to the sample of countries included in the regression. I construct the samples according to the national statistical concept that has been used to measure the size of the primate cities. In regression 3, I restrict the previous sample of 23 large countries with 4 cities above 750,000 inhabitants to those 19 countries that use urban agglomeration as the statistical concept. In regression 4, I also use countries that use metropolitan area as the statistical concept. In regression 5, I use all 44 countries that use urban agglomeration as the statistical concept, and in regression 6, I add those 10 countries that report data for metropolitan areas.<sup>13</sup> In regression 7, I use the full set of 70 countries, including those that report city proper data only.

Due to the different statistical concepts, the true degree of urban concentration is measured with an error that is increasing with the sample size. OLS estimates remain unbiased and consistent for measurement errors in the dependent variable, but they are not efficient. Therefore, lower levels of statistical significance are to be expected for the larger samples that use primacy measures on the basis of different statistical concepts.

Nevertheless, I find empirical evidence for the bell shaped relationship between urban concentration and GDP in all five regressions. The coefficients of *NRGDPC* vary between

12. The test augments the original regression by a matrix of test variables and tests the  $H_0$  that the elements of the coefficient vector of these variables are jointly zero. The additional variables can be the original independent variables raised to the 2nd, 3rd, 4th ... power, depending on the degrees of freedom remaining in the regression. This makes the test also powerful in detecting non-linearities and a wrong functional form. The drawback of the test is that it cannot discriminate between omitted variables and wrong functional form. See Zietz (1988) for a discussion of several adequacy and misspecification tests.

13. The 44 and 54 country samples are marked in Table A1 in the appendix.

## **JOURNAL OF ECONOMIC DEVELOPMENT**

0.1139 and 0.0488. The coefficients of *NRGDPC2* vary between -0.0021 and -0.0009. Both decrease in absolute values if the sample gets larger. The coefficients indicate that a maximum level of urban concentration is reached at RGDP levels between 11,371 and 13,062 US\$.

The bell shaped curve is shown in Figure 1 and Figure 2. Figure 1 is based on Regression 2 with 23 observations from large countries, Figure 2 is based on Regression 5 with all 44 observations that use urban agglomeration as the statistical concept. The horizontal axis shows PPP based GDP per capita. The vertical axis shows primacy net of the conditional variables. The figures display the individual data points and the estimated regression line. It can be seen that conditional primacy is first increasing and then decreasing with GDP per capita.

**Figure 1 The Bell Shaped Curve of the 23 Largest Countries**

**Figure 2 The Bell Shaped Curve of 44 Countries**

## JUNIUS: PRIMACY AND ECONOMIC DEVELOPMENT

As expected, the dummy variables for colonies in 1815 and colonies in 1950 increase urban concentration. It is remarkable that the coefficient of *NLAND* is significant at the 1 percent level in all regressions 2-7. *ROADLAND* enters all equations in that it is included with a negative coefficient. This indicates that a better transportation infrastructure, which means lower transport costs, mainly benefits non-core regions. Thus, the negative effect of a higher degree of competition is more than offset by the positive effect of a better access of the periphery to the core-market.

In regression 8-10, I use the original sample of 23 countries to test whether the results are sensitive to the measurement of urban concentration. I.e., I use *PRIMA2*, *PRIMA3* and *PRIMA4* as dependent variables. The coefficients and the statistical fit of regressions with different dependent variables cannot be compared. However, it can be seen that the existence of the bell shaped curve is robust for different primacy measures, because for all three regressions, I find that the coefficients of *NRGDP* and *NRGDP2* have the expected sign and are statistically significant.<sup>14</sup>

In regression 11 and 12, I test the hypothesis that countries that industrialized and, thus, urbanized relatively early are less concentrated today. This would reflect that industrial location decisions in the last century were dictated by natural endowments to a larger degree than nowadays. High costs of internal transportation led to the development of, for instance, steel production close to natural resources and transportation nodes in the “Ruhr area” in Germany and in the “Great Lake region” in the US. Consequently, early developers set up resource intensive industrial production not necessarily close to existing centers.<sup>15</sup> If so, industrialization might have led to population dispersion. Countries that industrialized relatively late could rely on more advanced transportation technologies. Therefore, they could set up industrial production close to the existing centers and benefit from the availability of a larger number of workers, intermediate goods suppliers and final demand. In this case, industrialization would increase urban concentration. I test this hypothesis adding *INDUSTRY1900* in regression 11 and *GDP1913* in regression 12 as measures of early development. As expected, both the share of the industrial sector as well as the per capita GDP level in 1913 have negative coefficients. However, they have large standard errors. The presence of the proposed curve is not destroyed by the inclusion of these variables.

Regression 13 and 14 further investigate the role of ancient conditions on current urban concentration. I exclude *DENSE* and *ROADLAND* from these regressions and include *DENSE1919*, *NTEL1919* and *RAIL1919*. The density of the telegraph system and the railroad system indicate the cost of overcoming distance. Population density is used as before. Due to a correlation coefficient of 85 percent between *DENSE1919* and *RAIL1919* individual effects of these variables cannot be distinguished and, hence, they are not included together in the equations. The results show that all variables enter the equation with a negative coefficient.

14. This confirms e.g., Rosen and Resnick (1980) who find a high correlation between different primacy measures. In my sample of 70 countries, the correlation between *PRIMA1* and *PRIMA2*, *PRIMA3* and *PRIMA4* is 97 percent, 93 percent and 89 percent, respectively. In my sample of 23 countries, that is used in regressions No. 1, 2, 8, 9, 10 the correlation is 95 percent, 91 percent and 88 percent, respectively.

15. For instance, Moky (1995) points out that it is striking that “neither Brussels nor Paris nor Berlin nor Amsterdam, nor any other major capital city in Europe, became a center of modern industry”.

Thus, early conditions seem to matter, again leaving intact the previously established bell shaped curve.<sup>16</sup>

Regression 15-17 explore whether historic patterns of urban concentration can add to the explanation of different degrees of urban concentration nowadays. This might correct for some idiosyncratic or genuine differences between countries resulting from a historic accident or geographic, climatic or endowment differences.<sup>17</sup> In turn, this might improve the significance of the other economic and conditional variables. However, it is also possible that adding historic degrees of concentration renders all other variables insignificant. In this case, current patterns would be entirely determined by historic patterns. Population growth and further urbanization would then only have resulted in a proportional growth of all cities. Degrees of concentration would then be determined by early historic accidents, climatic geographic or endowment differences. It seems reasonable that much of today's concentration can be explained by degrees of concentration in the 1980s and 1970s and may be even in the 1960s and 1950s, but it is unclear how long the effects take to peter out, if at all.

Therefore, I include, as further independent variables, measures of urban concentration at three points in time. In regression 15, I use the primacy ratio in 1800 (*HISTO1800*). The same ratio is calculated for 1850 (*HISTO1850*) and used in regression 16. In regression 17, I use the percentage of the largest in the three largest cities (*JEFFRATIO*), which is available for 41 countries in the 1930s from Jefferson (1939). As geographic and demographic differences of countries are reflected in the historic measures of urban concentration, only the political and economic explanatory variables are included in the regressions.

The coefficients of the primacy ratios of 1800 and 1850 in regressions 15 and 16 are significant at the 1 percent level. However, the *NRGDPC* and *NRGDPC2* coefficients are losing its significance if the historic variables are included in this sample. Whether this is due to the small sample size of 22 countries, the different statistical concepts of measuring city size in this sample and the evolving measurement errors in the dependent variable or the non-existence of the bell shaped curve in this sample remains unclear. Also the F-test and the RESET-test point to a misspecification if *HISTO1850* or *HISTO1800* are included and the previous explanatory variables are excluded in the sample. Using a sample of 41 countries in regression 17, the *JEFFRATIO* turns out to be significant at the 1 percent level. The inclusion of this ratio does not touch the existence of the bell shaped curve and the coefficient of *COLONY1815*. Apparently it reduces standard errors as indicated by the exceptionally high significance levels in this regression. Therefore, future research should continue in this line and better identify the role of path dependencies, and how long an impact historic patterns of urban concentration have on current patterns of urban concentration.

16. Note that the high possibility of having omitted relevant variables in regressions 13 and 14, as indicated by the F-statistic, is most likely due to a 85 percent correlation between *OPEN* and *NRGDPC*. Therefore, *OPEN* would indicate some of the effects that actually should be assigned to *NRGDPC*.

17. An example for such a historic accident is the dissolution of the Austrian-Hungarian empire which left a huge capital with a relatively small hinterland. This results in a very high primacy ratio for Austria.

## **V. Conclusion**

In this paper, I tested whether the relationship of primacy and economic development takes the form of a bell shaped curve, where concentration first increases and then decreases in the course of economic development. The alternative hypothesis was that spatial concentration is unaffected by economic development such that cities grow in a parallel way and primacy ratios remain constant. I find evidence for the first hypothesis, using different samples of countries. The bell shaped relationship is conditional on the size of the land area, population density and the density of the transportation system. I also find that countries with a long independent urban history have lower degrees of urban concentration today and countries with a relatively recent colonial past have higher degrees of urban concentration. Openness and political and economic freedom add little to the explanation of urban concentration. In none of the regressions, their coefficients are significant at the 10 percent level. This might be due to the measurement of this variables. Future empirical research should better identify the relationship of these variables and urban concentration. Historic values of population density and the extent of the transportation system also add to the explanation of urban concentration. This indicates that some determinants can have long lasting effects. 19th century degrees of urban concentration were found to have an impact on current patterns of concentration. The degree of concentration of the 1930s also improves the explanation of current concentration significantly.

Understanding the determinants of urban concentration and transition is important for being confident in future projections of urban growth and regional inequalities. These in turn, are essential to formulating proper regional, social and economic policies. The appeal of the bell shaped curve is also rooted in the hope that some problems especially developing countries have with regional inequalities and the excessive growth of their primate cities will vanish in the course of further development. This paper yields some evidence from cross-country regressions that support this hope.

Appendix

Table A1 Country List

		RGDPC	Primacy			RGDPC	Primacy
Algeria	*	2957	0.2353	Kenya		1080	0.2756
Argentina	*#	5532	0.3793	Korea	#	8271	0.3337
Australia	#+	17517	0.2452	Malaysia		5997	0.1275
Austria	*	15560	0.4814	Mali	*	714	0.3366
Bangladesh	+	1641	0.3462	Mexico	#+	6896	0.2459
Belgium	*	16533	0.1195	Morocco	*	2554	0.2509
Benin		1128	0.3989	Netherlands	*	16096	0.0794
Bolivia		1890	0.2815	New Zealand	*	14591	0.3078
Brazil	*#	4792	0.1340	Niger	*	563	0.3804
Bulgaria		7529	0.2157	Nigeria	*	1117	0.2287
Burkina Faso		608	0.3941	Norway	*	16345	0.2221
Cameroon		1249	0.2155	Pakistan	*#	1661	0.2205
Canada	*#	20752	0.1771	Paraguay	+	2496	0.2984
Chile	+	5279	0.4160	Peru	+	2603	0.4297
China	*#	1536	0.0452	Philippines	*	2112	0.2686
Colombia	*#	3902	0.2080	Poland	*#	4564	0.1448
Cote d'Ivoire	*	1372	0.4482	Portugal	*	9005	0.5015
Czech Republic	*	5066	0.990	Romania	*	2656	0.1697
Denmark	+	17217	0.3086	Russian Federation	*#	8780	0.0836
Ecuador	*	3163	0.2653	South Africa	*#	3886	0.1258
Egypt, Arab Rep. of	*	2153	0.3696	Spain	+	11765	0.1409
Finland	*	17080	0.2848	Sri Lanka		2468	0.1671
France	*#	16956	0.2264	Sweden	*	18024	0.2095
Germany	*#	18235	0.0938	Switzerland	*	20729	0.2031
Ghana	+	1101	0.2751	Syrian Arab Rep.	*	4714	0.2888
Greece	*	8203	0.5449	Thailand		4270	0.5708
Guatemala		2535	0.2324	Tunisia	*	3392	0.3930
Hungary		6430	0.3134	Turkey	*#	4489	0.1905
India	*#	1505	0.0567	Uganda	*	625	0.4019
Indonesia	*#	2323	0.1697	United Kingdom	*#	15741	0.1434
Iran, Islamic Rep. of	*#	3577	0.1914	United States	*#	21827	0.0854
Ireland	*	11273	0.4596	Uruguay		5536	0.4679
Italy	*#	15309	0.1210	Venezuela	#+	6859	0.1573
Japan	*#	17625	0.2623	Zaire		471	0.3284
Jordan		3774	0.4284	Zambia		799	0.2860

Notes: RGDPC denotes PPP-adjusted GDP per capita levels. Primacy denotes the share of the largest city in total urban population. Both values are for 1990. - #Country has at least 4 cities with more than 750,000 inhabitants (23 countries). - \*Country reports urban agglomeration data (44 countries). - +Country reports metropolitan area data (10 countries).

Source: UN (1995a), World Bank (1992).



JUNIUS: PRIMACY AND ECONOMIC DEVELOPMENT

**Table A2 Regression Results**

Regression Dependent variable	Basic results		Different samples		
	1	2	3	4	5
	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>
<i>NRCGDCP</i>	0.1120** (0.0387)	0.1139*** (0.0388)	0.0829** (0.0346)	0.0833** (0.0289)	0.0676** (0.0287)
<i>NRCGDPC2</i>	-0.0021** (0.0008)	-0.0021*** (0.0007)	-0.0015** (0.0007)	-0.0015** (0.0006)	-0.0013** (0.0006)
<i>NLAND</i>	-0.0027*** (0.0008)	-0.0023*** (0.0006)	-0.0028*** (0.0005)	-0.0027*** (0.0005)	-0.0035*** (0.0005)
<i>DENSE</i>	-0.0006 (0.0020)		-0.0032 (0.0019)	-0.0030 (0.0017)	-0.0037*** (0.0010)
<i>ROADLAND</i>	-0.1996 (0.3997)	-0.2229 (0.2854)			
<i>COLONY1815</i>	0.4624 (0.2746)	0.4224* (0.2365)	0.3644 (0.2186)	0.4097* (0.1944)	0.3641** (0.1767)
<i>COLONY1950</i>					0.4645* (0.2533)
<i>POLITICS</i>	0.2877 (0.2937)				
<i>OPEN</i>	-0.1809 (0.3649)				
<b>Historic variables</b>					
<i>INDUSTRY1900</i>					
<i>GDPC1913</i>					
<i>DENSE1919</i>					
<i>NTEL1919</i>					
<i>RAIL1919</i>					
<i>JEFFRATIO</i>					
<i>HISTO1850</i>					
<i>HISTO1800</i>					
$\bar{R}^2$	0.20	0.30	0.38	0.42	0.53
SEE	0.56	0.53	0.49	0.46	0.53
No. of observations	23	23	19	22	44
Jarque-Bera test for normality of residuals <sup>a</sup>	1.08 (0.58)	1.35 (0.51)	0.38 (0.82)	0.25 (0.88)	0.40 (0.50)
White test for hetero-scedasticity <sup>a</sup>	2.09 (0.13)	1.06 (0.44)	2.32 (0.11)	3.21 (0.03)	0.76 (0.65)
Ramsey reset test (3) <sup>a</sup>	1.86 (0.20)	0.52 (0.68)	2.48 (0.12)	1.93 (0.17)	0.40 (0.76)
F-test for omitted variables <sup>a</sup>	-	0.27 (0.84)	0.88 (0.48)	0.91 (0.46)	0.45 (0.72)

Standard deviations in parenthesis: \*Significant at 10 percent level; \*\*significant at 5 percent level; \*\*\*significant at 1 percent level. - SEE = Standard error of estimates. - <sup>a</sup>( )Marginal probability values in parenthesis. - Constants are not reported.

Table A2 (Continued)

Regression Dependent variable	Different samples		Different measures of urban concentration		
	6	7	8	9	10
	<i>PRIMA</i> 1	<i>PRIMA</i> 1	<i>PRIMA</i> 2	<i>PRIMA</i> 3	<i>PRIMA</i> 4
<i>NRCGDCP</i>	0.0658** (0.0312)	0.0488 (0.0340)	0.1039*** (0.0351)	0.0985*** (0.0319)	0.1008*** (0.0331)
<i>NRCGDPC2</i>	-0.0012* (0.0006)	-0.0009 (0.0007)	-0.0018** (0.0006)	-0.0017** (0.0006)	-0.0017** (0.0006)
<i>NLAND</i>	-0.0033*** (0.0006)	-0.0035*** (0.0006)	-0.0019*** (0.0006)	-0.0022*** (0.0006)	-0.0023*** (0.0006)
<i>DENSE</i>	-0.0011 (0.0010)	-0.0008 (0.0008)			
<i>ROADLAND</i>	-0.2003 (0.1323)	-0.2245* (0.1168)	-0.2077 (0.2859)	-0.2723 (0.2687)	-0.3071 (0.2651)
<i>COLONY1815</i>	0.4888*** (0.1662)	0.3802** (0.1658)	0.5410** (0.1985)	0.5326*** (0.1805)	0.5617*** (0.1841)
<i>COLONY1950</i>	0.4885* (0.2614)	0.2242 (0.2625)			
<i>POLITICS</i>					
<i>OPEN</i>					
<b>Historic variables</b>					
<i>INDUSTRY1900</i>					
<i>GDPC1913</i>					
<i>DENSE1919</i>					
<i>NTEL1919</i>					
<i>RAIL1919</i>					
<i>JEFFRATIO</i>					
<i>HISTO1850</i>					
<i>HISTO1800</i>					
$\bar{R}^2$	0.41	0.31	0.30	0.37	0.41
SEE	0.57	0.60	0.49	0.45	0.45
No. of observations	54	70	23	23	23
Jarque-Bera test for normality of residuals <sup>a</sup>	0.79 (0.67)	0.06 (0.97)	1.02 (0.60)	0.84 (0.66)	0.96 (0.62)
White test for hetero-scedasticity <sup>a</sup>	1.13 (0.36)	1.27 (0.26)	2.08 (0.11)	1.91 (0.14)	1.64 (0.20)
Ramsey reset test (3) <sup>a</sup>	0.49 (0.69)	1.45 (0.24)	1.30 (0.31)	1.45 (0.27)	1.60 (0.23)
F-test for omitted variables <sup>a</sup>	0.06 (0.94)	0.77 (0.47)	0.13 (0.94)	0.16 (0.92)	0.20 (0.90)

Standard deviations in parenthesis: \*Significant at 10 percent level; \*\*significant at 5 percent level; \*\*\*significant at 1 percent level. - SEE = Standard error of estimates. - <sup>a</sup>( )Marginal probability values in parenthesis. - Constants are not reported.

JUNIUS: PRIMACY AND ECONOMIC DEVELOPMENT

**Table A2 (Continued)**

Regression Dependent variable	Historic indicators		Historic infrastructure		Historic urban concentration		
	11	12	13	14	15	16	17
	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>	<i>PRIMA1</i>
<i>NRCGDCP</i>	0.0897* (0.0452)	0.0905** (0.0336)	0.0835** (0.0320)	0.0711** (0.0337)	0.0229 (0.0624)	0.0225 (0.0650)	0.0847** (0.0381)
<i>NRCGDCP2</i>	-0.0017* (0.0008)	-0.0016*** (0.0006)	-0.0017** (0.0006)	-0.0014** (0.0007)	-0.0008 (0.0013)	-0.0007 (0.0014)	-0.0017** (0.0008)
<i>NLAND</i>	-0.0032*** (0.0005)	-0.0033*** (0.0005)	-0.0035*** (0.0005)	-0.0034*** (0.0006)			
<i>DENSE</i>	-0.0041*** (0.0012)						
<i>ROADLAND</i>		-0.3021*** (0.1025)					
<i>COLONY1815</i>	0.2539 (0.2012)	0.5266* (0.2182)	0.4162** (0.1758)	0.5116*** (0.1861)	0.7835** (0.3617)	0.6693* (0.3532)	0.6213*** (0.2248)
<i>COLONY1950</i>							
<i>POLITICS</i>							
<i>OPEN</i>							
<b>Historic variables</b>							
<i>INDUSTRY1900</i>	-0.8028 (0.8303)						
<i>GDPC1913</i>		-0.0001 (0.0001)					
<i>DENSE1919</i>			-0.0002*** (0.0001)				
<i>NTEL1919</i>			-0.0018 (0.0013)	-0.0022* (0.0013)			
<i>RAIL1919</i>				-0.0003** (0.0001)			
<i>JEFFRATIO</i>							3.0756*** (0.7086)
<i>HISTO1850</i>						2.7574*** (0.6449)	
<i>HISTO1800</i>					2.3547*** (0.6391)		
$\bar{R}^2$	0.41	0.30	0.44	0.38	0.32	0.40	0.38
SEE	0.61	0.67	0.58	0.61	0.66	0.62	0.60
No. of observations	27	44	45	45	22	22	41
Jarque-Bera test for normality of residuals <sup>a</sup>	0.76 (0.68)	0.35 (0.84)	0.73 (0.70)	0.48 (0.79)	1.00 (0.61)	0.87 (0.65)	1.29 (0.52)
White test for hetero-scedasticity <sup>a</sup>	1.36 (0.28)	0.50 (0.88)	0.44 (0.91)	0.50 (0.88)	0.46 (0.82)	1.04 (0.44)	0.39 (0.88)
Ramsey reset test (3) <sup>a</sup>	0.50 (0.68)	0.21 (0.89)	0.35 (0.79)	0.33 (0.80)	2.17 (0.14)	1.18 (0.35)	0.39 (0.76)
F-test for omitted variables <sup>a</sup>	0.78 (0.52)	1.50 (0.23)	3.65 (0.04)	2.41 (0.10)	3.40 (0.05)	2.84 (0.08)	3.15 (0.03)

Standard deviations in parenthesis: \*Significant at 10 percent level; \*\*significant at 5 percent level; \*\*\*significant at 1 percent level. - SEE = Standard error of estimates. - <sup>a</sup>( ) Marginal probability values in parenthesis. - Constants are not reported.

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